

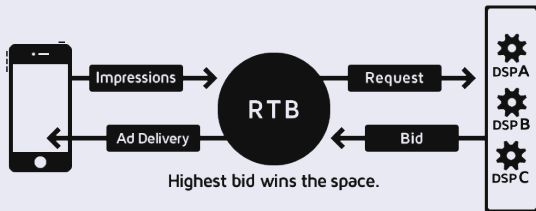
# Adaptive On-line Learning for First Bid Prediction and Reserve Price Optimisation

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# Real Time Bidding

- Programmatic instantaneous auction to sell/buy Ads
- Two main actors: Demand-side platforms (DSP) supporting buyers and Supply-side platforms (SSP) supporting publishers

## Impressions and advertisement



## Second bid

- For a given auction:  $B = \{b_1 > b_2 > \dots b_n\}$ , set of sorted bids.  $r$ , floor price
- Revenue  $\pi$  of publisher is:

$$\pi(b_1, b_2, r) = \begin{cases} b_2 & \text{if } b_2 \geq r \\ r & \text{if } b_1 \geq r > b_2 \\ 0 & \text{if } r > b_1 \end{cases} = \max(b_2, r) 1_{r < b_1}$$

## Soft Floor

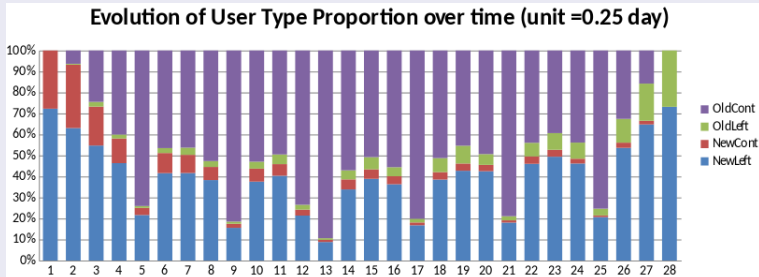
Some publishers consider also a soft floor,  $s$

$$= \pi(b_1, b_2, r, s) = \max(b_2, r, \min(b_1, s)) 1_{r < b_1}$$

# Estimation of the highest bid

## Challenges

- Unobserved (censored) data when  $B1 < r$
- Data: time, user, tag, highest bid, paid price, reserve price
- No external data( impression, user localisation, user device.... )
- "New users", in many cases there is not historical information

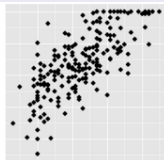
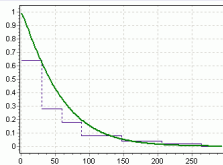


# Dealing with censored data

## Reliability (survival) function <sup>1</sup>

Probability of an event  $T$  not occurring yet at time  $T$

$$S(t) = P(T > t)$$



Kaplan Meier Estimator of the survival function  $S(t)$

$$S(t) = \prod_{t_i < t} \frac{n_i - d_i}{n_i}$$

Aalen Additive Model [Aalen, 1980] - linear model with evidence  $x$

$$\lambda(t) = b_0(t) + b_1(t)x_1 + b_2(t)x_2 + \dots + b_n(t)x_n \text{ with } \lambda(t) = S'(t)/S(t)$$

<sup>1</sup>Kaplan, E. L.; Meier, P. (1958). "Nonparametric estimation from incomplete observations".

# Tobit Model - Dealing with censored data

- Estimate a latent variable  $y_i^*$  given independent  $x_i$
- Observed  $y_i$

$$y_i = \begin{cases} y_i^* & \text{if } y_i^* \geq \tau_i \\ \tau_i & \text{if } y_i^* < \tau_i \end{cases}$$

$$y_i^* = \beta x_i + u_i, u_i \sim N(0, \sigma^2)$$

- log-Likelihood of  $\beta$   
 $\log \mathcal{L} = \sum_i \log\left(\frac{1}{\sigma} \phi\left(\frac{y_i - x_i \beta}{\sigma}\right)\right) \mathbf{1}_{y_i^* \geq \tau_i} + \log\left(\Phi\left(\frac{x_i \beta - \tau}{\sigma}\right)\right) \mathbf{1}_{y_i^* < \tau_i}$
- Optimal parameters can be found by Newton-Raphson method

## Recursive Least Square - Linear Regression

$$Y_i^* = \beta X_i + u_i, u_i \sim N(0, \sigma^2)$$

- Update parameters  $\beta$  at each observation without requiring the previous set  $X_{1..N}$

$$S_N = X'_{1..N} X_{1..N}$$

$$S_{N+1} = S_N + X'_{N+1} X_{N+1}$$

$$\gamma_{N+1} = S_{N+1}^{-1} + X'_{N+1}$$

$$e = Y_{N+1} - X_{N+1} \hat{\beta}_N$$

$$\hat{\beta}_{N+1} = \hat{\beta}_N + \gamma_{N+1} e$$

- It cannot be directly applied when the dependent variable is censored

# Adaptive estimation of a truncated univariate gaussian variable

- predicting  $y \sim N(u, \sigma)$
- $\lambda(x) = \frac{\phi(x)}{1-\Phi(x)}$ ,  $\rho(x) = \frac{\phi(x)}{\Phi(x)}$ ,  $\gamma$  : *factor*

Prediction:  $\hat{y} = u$

Update:

If  $y$  is not censored :

$$S = \gamma * S + y$$

$$SS = \gamma * SS + (y - u)^2$$

else:

$$M1 = (u - \sigma * \rho((\tau - u)/\sigma))$$

$$S = \gamma * S + M1$$

$$M2 = \sigma^2 + u * M1 - \sigma * \tau * \rho((\tau - u)/\sigma)$$

$$SS = \gamma * SS + (M2 - 2 * u * M1 + u * u)$$

$$u = S/n$$

$$\sigma = \sqrt{SS/n}$$



# The Kalman Filter

- A filter to estimate the state of a discrete time dynamic linear system using noisy measurements

- The state of the system at time  $k$  is given by:

$$x_{k+1} = A_k x_k + B_k u_k + w_k \quad w_k \sim N(0, Q_k)$$

- A measurement:

$$y_k = C_k x_k + v_k \quad v_k \sim N(0, R_k)$$

- Conditional pdf of the state:

$$p(x_k | Y_k, U_{k-1}) \sim N(\hat{x}(k | k), P(k | k))$$

- Kalman Gain:

$$K(k+1) = P(k+1 | k) C'_{k+1} [C_{k+1} P(k+1 | k) C'_{k+1} + R]^{-1}$$

## State estimation

- Prediction step

$$\hat{x}(k+1 | k) = A_k \hat{x}(k | k) + B_k u_k$$

$$P(k+1 | k) = A_k P(k | k) A'_k + Q$$

- Filtering step

$$\hat{x}(k | k) = \hat{x}(k | k-1) + K(k) [y(k) - C_k \hat{x}_{k|k-1}]$$

$$P(k | k) = [I - K(k) C_k] P(k | k-1)$$

# Kalman Filter for Linear Regression

- Problem definition:  $\hat{\beta}_{k+1} = \hat{\beta}_k$   
 $y_k = z_k' \hat{\beta}_k + e_k, e_k \sim N(0, 1)$
- Kalman filter matrices:  $A = I, R = 0, Q = 1, C = z_k'$
- Kalman filter equations:  
$$K(k+1) = P(k+1 | k) z_{k+1}' [z_{k+1} P(k+1 | k) z_{k+1}' + 1]^{-1}$$
$$\hat{\beta}_{k+1} = \hat{\beta}_k + K(k+1) [y_k - z_k' \hat{\beta}_k]$$
$$P(k | k) = [I - K(k) z_k] P(k | k-1)$$

- Extension of Kalman Filter to treat truncated data
- Prediction step remains the same:
- Filtering step:

- $S(k | k - 1) = C_k P(k | k - 1) C_k' + R$

- If  $Y_k$  is not censored:

- $\hat{x}(k | k) = \hat{x}(k | k - 1) + K(k)[y(k) - C_k \hat{x}_{k|k-1}]$

- $P(k | k) = [I - K(k)C_k]P(k | k - 1) = P(k | k - 1) - K(k)S(k | k - 1)K'(k)$

- If  $Y_k$  is censored:

- $\hat{x}(k | k) = \hat{x}(k | k - 1) + K(k)[E(y(k) | y(k) \leq \tau) - C_k \hat{x}_{k|k-1}]$

- $P(k | k) = P(k | k - 1) - K(k)(S(k | k - 1) - \text{Var}(y(k) | y(k) \leq \tau))K'(k)$

- In the case of a univariate variable  $\sim N(m, \sigma)$  with  $m = C_k \hat{x}_{k|k-1}$

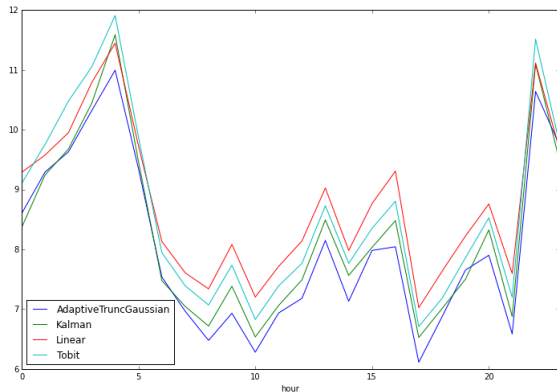
$$E(y(k) | y(k) \leq \tau) = m - \sigma \frac{\phi(\frac{\tau-m}{\sigma})}{\Phi(\frac{\tau-m}{\sigma})}$$

$$\text{Var}(y(k) | y(k) \leq \tau) = m^2 - 2m\sigma \frac{\phi(\frac{\tau-m}{\sigma})}{\Phi(\frac{\tau-m}{\sigma})} + \sigma^2(1 - \frac{\tau-m}{\sigma})$$

# Some results in Predicting the Highest bid(MAE)

- Features: Previous Highest Price(B1), Previous B1 for same user, Previous B1 for same tag

Mean average error per hour:



# Revenue Maximization

- The revenue function is asymmetric: underestimating the highest bid is better than overestimating it
- Maximizing the revenue is more than predicting the highest bid

In a second bid setting.

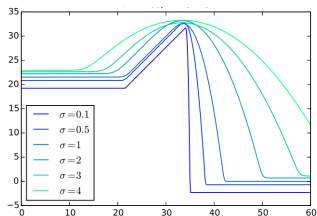
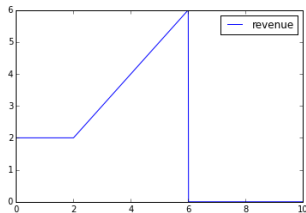
Modeling the reserve price as  $f(x_i, w_i) = w'x_i$

Objective: find  $w$  that maximize the revenue:

$$w^* = \operatorname{argmax} \sum_i^N R(f(x_i, w_i), B1_i, B2_i)$$

Smoothing the revenue:  $y_i \sim N(f(x_i, w_i), \sigma^2)$

$$L(w) = \sum_i^N \log E[\exp(R(y_i, B1_i, B2_i))]$$



## Application of latent variables [Rudolph 2014]

- Variable  $z$  indicates satisfiability on the auction

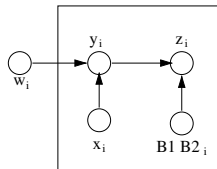
$$p(z_i = 1 \mid y_i, B1_i, B2_i) = \pi(y_i, B1_i, B2_i)$$

$$\text{where } \pi(y_i, B1_i, B2_i) = \exp(-(B1_i - R(y_i, B1_i, B2_i)))$$

The proposed model is:

$$y_i \mid w, x_i \sim N(f(x_i, w), \sigma^2)$$

$$z_i \mid y_i, B1_i, B2_i \sim \text{Bernoulli}(\pi(y_i, B1_i, B2_i))$$



The smoothed revenue can be maximized by MAP estimates of  $w$ :

$$\text{logp}(w \mid z, x, B1, B2) = \sum_i^N (\text{log}E[\exp R(y_i, B1_i, B2_i)] - B1_i)$$

- Integrating Kalman filter into Revenue Maximization
- Considering non linear models
- Testing on new datasets
- Matrix factorization - Latent variables



Yuan, Shuai and Wang, Jun and Chen, Bowei and Mason, Peter and Seljan, Sam. An Empirical Study of Reserve Price Optimisation in Real-time Bidding. KDD 14.2014.



Maja R. Rudolph, Joseph G. Ellis, David M. Blei. Objective Variables for Probabilistic Revenue Maximization in Second-Price Auctions with Reserve. arXiv preprint arXiv:1506.07504.2015